

Statistical inference in transport-fragmentation equations

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- We consider (simple) particle systems \approx toy models for the evolution of cells or bacteria.
 - Each particle grows by ingesting a common nutrient.
 - After some time, each particle gives rise to two offsprings by cell division.
- We structure the model by state variables like **size, growth rate** and so on.
 - Deterministically, the density of structured state variables evolves according to a **fragmentation-transport PDE**.
 - Stochastically, the particles evolve according to a **PDMP** that evolves along a **branching tree**.

Toy model : size-structured populations

- $n(t, x)$: density of cells of size x .
- Parameter of interest : **Division rate** $B(x)$.
- 1 cell of size x gives birth to 2 cells of size $x/2$.
- The growth of the cell size by nutrient uptake is given by a **growth rate** $g(x) = \tau x$ (for simplicity).

Structured populations (cont.)

- **Transport-fragmentation** equation

$$\partial_t n(t, x) + \partial_x (\tau x n(t, x)) + B(x) n(t, x) = 4B(2x) n(t, 2x)$$

with $n(t, x = 0) = 0$, $t > 0$ and $n(0, x) = n^{(0)}(x)$, $x \geq 0$.

- obtained by mass conservation law :

- LHS : density evolution + growth by nutrient + division of cells of size x .
- RHS : division of cells of size $2x$.

- Several extensions...

Objectives

- Main goal : **estimate non-parametrically B** from genealogical data of a cell population of size N living on a binary tree.
- **Avoid an inverse problem** (cp. Doumic, Perthame, Zubelli, 2009 & Doumic, H., Reynaud, Rivoirard, 2012) thanks to **richer data set**.
- **Reconcile** the deterministic approach with a rigorous statistical analysis (relaxing the steady-state approximation).

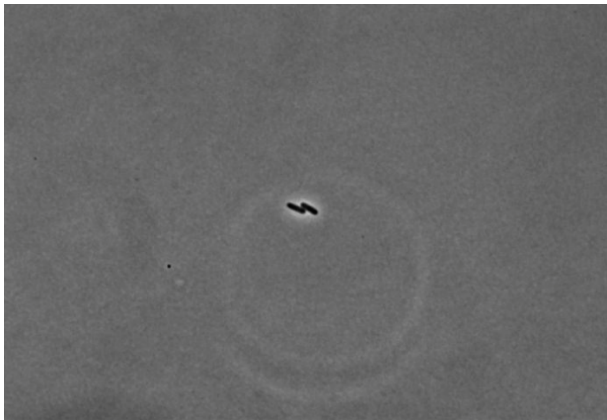


FIGURE : Evolution of a *E. Coli* population.

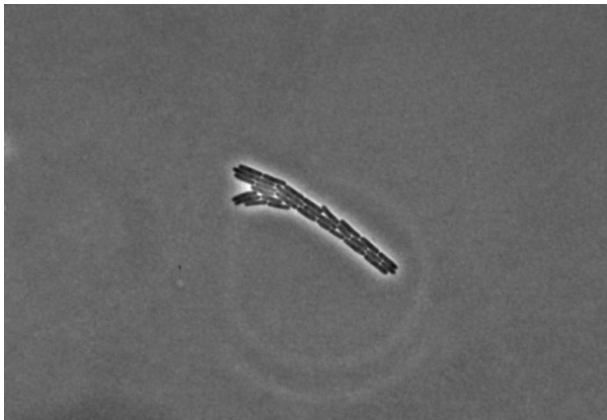


FIGURE : Evolution of a *E. Coli* population.

- Construct a stochastic model accounting for the **stochastic dependence structure on a tree** for which the empirical measure of N particles solves the fragmentation-transport equation (in a weak sense).
- Develop **appropriate statistical tools** to estimate B .
- Additional difficulty & goal : incorporate **growth variability** (each cell has a stochastic growth rate inherited from its parent).

Result 1

- We construct a Markov process on a binary tree

$$(X_t, V_t) \in \left(\bigcup_{k \geq 0} [0, \infty)^k \right)^2,$$

where X_t =size and V_t =growth rate of living cells at time t (inherited from their parent according to a kernel ρ).

- $n(t, \cdot) := \mathbb{E} \left[\sum_{i=1}^{\infty} \delta_{X_i(t), V_i(t)} \right]$ is a (weak)-solution of an **extension** of the transport-fragmentation equation :

$$\begin{aligned} & \partial_t n(t, x, \mathbf{v}) + \mathbf{v} \partial_x (x n(t, x, \mathbf{v})) + B(x) n(t, x, \mathbf{v}) \\ &= 4 \int \rho(\mathbf{v}', d\mathbf{v}) n(t, 2x, d\mathbf{v}'). \end{aligned}$$

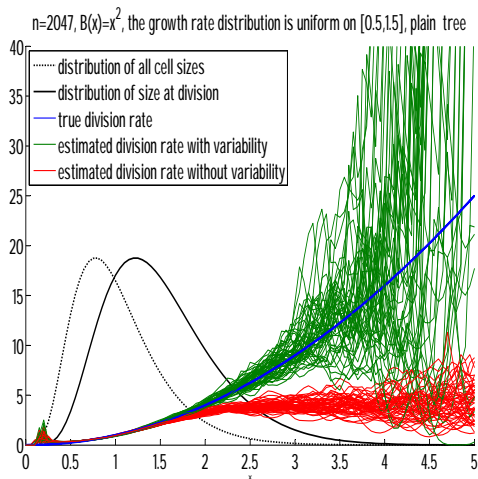
- The initial framework $g(x) = \tau x$ is retrieved as soon as $\rho(d\mathbf{v}) = \delta_{\tau}(d\mathbf{v})$

Result 2

- **Genealogical data** : we observe size+variability $(\xi_u, \tau_u)_{u \in \mathcal{U}_N}$, where \mathcal{U}_N is a (connected) subset of size N of the binary tree $\mathcal{U} = \cup_{k \geq 0} \{0, 1\}^k$.
- **Main result** : We can construct an estimator $(\hat{B}_N(x), x > 0)$ of the s -regular division rate $B(x)$ s.t.

$$\mathbb{E} [\|\hat{B}_N - B\|_{L_{\text{loc}}^2}^2]^{1/2} \lesssim (\log N) N^{-s/(2s+1)}$$

Numerical implementation and effect of variability



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FIGURE : Simulated data

Comparison with the inverse problem approach

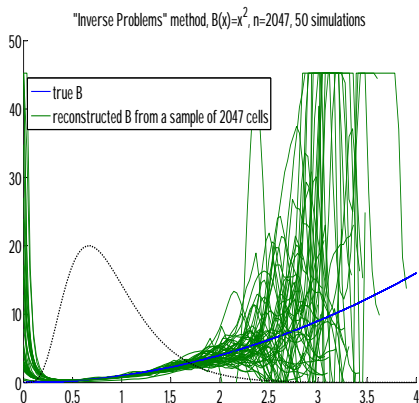


FIGURE : Exploration on simulated data via the global approach (inverse problem), $N \approx 3000$.

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Numerical implementation

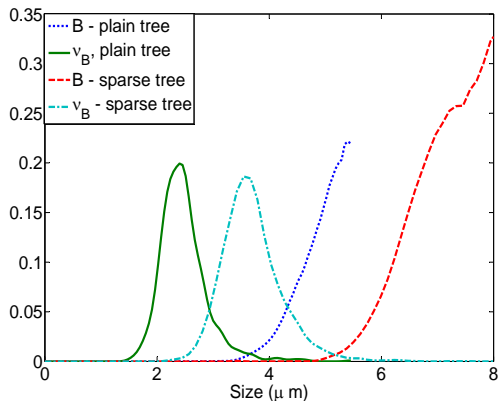


FIGURE : Exploration on real-data. Sparse tree, $N \approx 3000$.

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- Given a pair ξ_{u^-}, ζ_{u^-} and ξ_u , we can identify τ_{u^-} through

$$2\xi_u = \xi_{u^-} e^{\tau_{u^-} \zeta_{u^-}}.$$

- We have

$$\mathbb{P}(\zeta_u \in [t, t + dt] | \zeta_u \geq t, \xi_u = x) = B(xe^{\tau t}) dt$$

from which we obtain the **density of the lifetime** ζ_{u^-} (lifetime of the parent u^-) conditional on $\xi_{u^-} = x$ and $\tau_{u^-} = v$:

$$t \rightsquigarrow B(xe^{vt}) \exp\left(-\int_0^t B(xe^{vs}) ds\right).$$

- Using $2\xi_u = \xi_{u^-} \exp(\tau_{u^-} \zeta_{u^-})$, we further infer

$$\begin{aligned} & \mathbb{P}(\xi_u \in dx' \mid \xi_{u^-} = x, \tau_{u^-} = v) \\ &= \frac{B(2x')}{vx'} \mathbf{1}_{\{x' \geq x/2\}} \exp\left(-\int_{x/2}^{x'} \frac{B(2s)}{vs} ds\right) dx'. \end{aligned}$$

- We obtain a simple and explicit representation for the transition kernel on $\mathcal{S} = [0, \infty) \times \mathcal{E}$,

$$\mathcal{P}_B(\mathbf{x}, d\mathbf{x}') = \mathcal{P}_B((x, v), x', dv') dx'$$

The explicit transition

- The formula is given by

$$\begin{aligned} & \mathcal{P}_B((x, v), x', dv') dx' \\ &= \frac{B(2x')}{vx'} \mathbf{1}_{\{x' \geq x/2\}} \exp\left(-\int_{x/2}^{x'} \frac{B(2s)}{vs} ds\right) \rho(v, dv'). \end{aligned}$$

- $\rho(a, da')$: appropriate Markov kernel on \mathcal{E} .
- Under appropriate conditions on B , the Markov chain on $\mathcal{S} = [0, \infty) \times \mathcal{E}$ is geometrically ergodic. (It is however not reversible.)

Identifying B through the invariant measure

- Under appropriate assumptions, we have existence (and uniqueness) of an **invariant measure** on \mathcal{S}

$$\nu_B(d\mathbf{x}) = \nu_B(x, dv)dx$$

i.e. such that $\nu_B \mathcal{P}_B = \nu_B$.

- More precisely, we have a **contraction property**

$$\sup_{|g| \leq V} |\mathcal{P}_B^k g(\mathbf{x}) - \int_{\mathcal{S}} g(\mathbf{z}) \nu_B(d\mathbf{z})| \leq RV(\mathbf{x}) \gamma^k$$

for some $\gamma < 1$ locally uniformly in B for an appropriate Lyapunov function V .

Identifying B through the invariant measure

$$\begin{aligned} & \nu_B(y, dv') \\ &= \int_{\mathcal{S}} \nu_B(x, dv) dx \mathcal{P}_B((x, v), y, dv') \\ &= \frac{B(2y)}{y} \int_{\mathcal{E}} \int_0^{2y} \nu_B(x, dv) dx \exp\left(-\int_{x/2}^y \frac{B(2s)}{vs} ds\right) \frac{\rho(v, dv')}{v}. \end{aligned}$$

“Survival analysis trick”

$$\exp\left(-\int_{x/2}^y \frac{B(2s)}{vs} ds\right) = \int_y^\infty \frac{B(2s)}{vs} \exp\left(-\int_{x/2}^v \frac{B(2s')}{vs'} ds'\right) ds$$

and \mathcal{P}_B is involved in the RHS again...

Identifying B through the invariant measure

- We obtain

$$\begin{aligned}\nu_B(y, dv') &= \frac{B(2y)}{y} \int_{\mathcal{E}} \int_0^{2y} \nu_B(x, dv) dx \\ &\quad \int_y^\infty \frac{B(2s)}{vs} \exp\left(-\int_{x/2}^s \frac{B(2s')}{vs'} ds'\right) ds \frac{\rho(v, dv')}{v} \\ &= \frac{B(2y)}{y} \int_S \int_{[0, \infty)} \mathbf{1}_{\{x \leq 2y, s \geq y\}} v^{-1} \\ &\quad \nu_B(x, dv) dx \mathcal{P}_B((x, v), s, dv') ds.\end{aligned}$$

and integrate in dv' .

Key representation

- We obtain the **key representation**

$$\nu_B(y) = \frac{B(2y)}{y} \mathbb{E}_{\nu_B} \left[\frac{1}{\tau_{u^-}} \mathbf{1}_{\{\xi_u^- \leq 2y, \xi_u \geq y\}} \right].$$

- We conclude

$$B(y) = \frac{y}{2} \frac{\nu_B(y/2)}{\mathbb{E}_{\nu_B} \left[\frac{1}{\tau_{u^-}} \mathbf{1}_{\{\xi_u^- \leq y, \xi_u \geq y/2\}} \right]}.$$

- Introduce a kernel function

$K : [0, \infty) \rightarrow \mathbb{R}$, $\int_{[0, \infty)} K(y) dy = 1$ and set

$K_h(y) = h^{-1} K(h^{-1}y)$ for $y \in [0, \infty)$ and $h > 0$.

- Final estimator

$$\hat{B}_n(y) = \frac{y}{2} \frac{n^{-1} \sum_{u \in \mathcal{U}_n} K_h(\xi_u - y/2)}{n^{-1} \sum_{u \in \mathcal{U}_n} \frac{1}{\tau_{u^-}} \mathbf{1}_{\{\xi_{u^-} \leq y, \xi_u \geq y/2\}}} \vee \varpi$$

- The estimator $\hat{B}_n(y)$ is specified by K , **the bandwidth** h and the threshold ϖ .

Some references

- Doumic, M., H. M., Reynaud-Bouret, P. and Rivoirard, V. (2012) *Nonparametric estimation of the division rate of a size-structured population*. SIAM Journal on Numerical Analysis. **50**, 925-950
- Doumic, M., H.,M., Krell, N. and Robert, L. (2012) *Statistical estimation of a growth-fragmentation model observed on a genealogical tree*. arXiv :1210.3240
- Doumic, M., Perthame, B. and Zubelli, J. (2009) *Numerical Solution of an Inverse Problem in Size-Structured Population Dynamics*. Inverse Problems, **25**, 25pp.